

# Kriging of financial term-structures

Areski Cousin  
ISFA, Université Lyon 1

Joint work with Hassan Maatouk and Didier Rullière

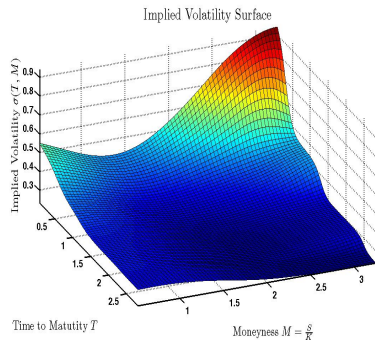
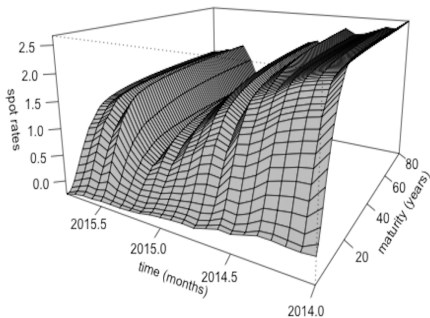
International Conference on Monte Carlo techniques

Paris, July 5th 2016



# Motivation

- Financial term-structures describes the evolution of some financial or economic quantities as a function of time horizon.
- **Examples** : interest-rates, bond yields, credit spreads, implied default probabilities, implied volatilities.
- **Applications** : valuation of financial and insurance products, risk management



## Several constraints have to be considered

- **Compatibility with market information** : at a given date  $t_0$ , the curve under construction  $T \rightarrow P(t_0, T)$  shall be compatible with observed prices of some reference products.
- **Arbitrage-free construction** : this translates into some specific shape properties such as positivity, monotonicity, convexity or bounds on the curve values
- **Additional conditions can be required** : minimum degree of smoothness, control of local convexity

# The term-structure construction problem

## 1) Compatibility with market information :

- At time  $t_0$ , we observe the market quotes  $S_1, \dots, S_n$  of  $n$  liquidly traded instruments
- The values of these products depend on the value of the curve at  $m$  input locations  $X = (\tau_1, \dots, \tau_m)$

The vector of output values  $P(t_0, X) := (P(t_0, \tau_1), \dots, P(t_0, \tau_m))^T$  satisfies a linear system of the form

$$A \cdot P(t_0, X) = \mathbf{b},$$

where

- $A$  is a  $n \times m$  real-valued matrix
- $\mathbf{b}$  is a  $n$ -dimensional column vector

$n < m \implies$  indirect and partial information on the curve values at  $\tau_1, \dots, \tau_m$

## 2) No-arbitrage assumption :

$T \rightarrow P(t_0, T)$  is typically a monotonic bounded function

We observe the quoted par rates  $S_i$  of an OIS with maturities  $T_i$ ,  $i = 1, \dots, n$

## 1) Compatibility with market quotes :

The curve  $T \rightarrow P(t_0, T)$  of **OIS discount factors** is such that

$$S_i \sum_{k=1}^{p_i-1} \delta_k P(t_0, t_k) + (S_i \delta_{p_i} + 1) P(t_0, T_i) = 1, \quad i = 1, \dots, n$$

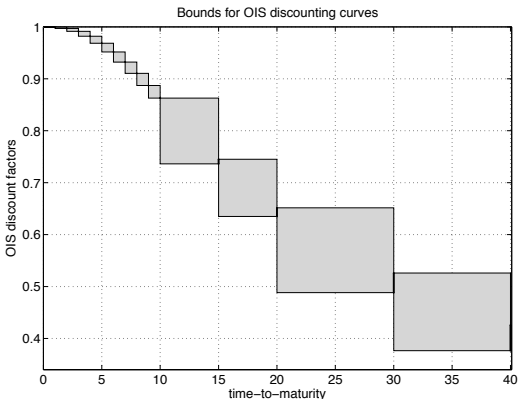
- $t_1 < \dots < t_{p_i} = T_i$  : fixed-leg payment dates (annual time grid)
- $\delta_k$  : year fraction of period  $(t_{k-1}, t_k)$

## 2) No-arbitrage assumption :

$T \rightarrow P(t_0, T)$  is a decreasing function such that  $P(t_0, t_0) = 1$

# Range of arbitrage-free OIS discount curves

- $n = 14$  liquidly traded maturities 1, 2, ..., 10, 15, 20, 30, 40 years.
- $m = 40$  points involved in the market-fit linear system
- No-arbitrage bounds on OIS discount factors



Input data : OIS swap rates as of May, 31st 2013.

Source : [Cousin and Niang \(2014\)](#)

We observe at time  $t_0$  the fair spreads  $S_i$  of a CDS with maturities  $T_i$ ,  
 $i = 1, \dots, n$

## 1) Compatibility with market quotes :

The curve  $T \rightarrow P(t_0, T)$  of (risk-neutral) **survival probabilities** is such that

$$S_i \sum_{k=1}^{P_i} \delta_k D(t_0, t_k) P(t_0, t_k) = -(1 - R) \int_{t_0}^{T_i} D(t_0, u) dP(t_0, u), \quad i = 1, \dots, n$$

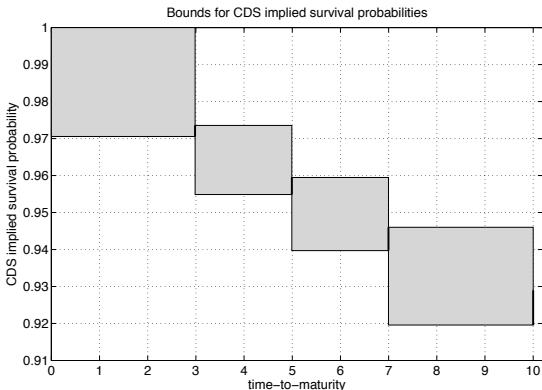
- $t_1 < \dots < t_p = T_i$  : trimestrial premium payment dates,  $\delta_k$  : year fraction of period  $(t_{k-1}, t_k)$
- $D(t_0, T)$  is the discount factor associated with maturity date  $T$
- $R$  : expected recovery rate of the reference entity

## 2) No-arbitrage assumption :

$T \rightarrow P(t_0, T)$  is a decreasing function such that  $P(t_0, t_0) = 1$

# Range of arbitrage-free CDS-implied survival functions

- $n = 4$  liquidly traded maturities 3, 5, 7, 10 years.
- $m = 40$  points involved in the market-fit linear system
- No-arbitrage bounds on the issuer implied survival distribution function



**Input data** : CDS spreads of AIG as of December 17, 2007,  $R = 40\%$ ,  
 $D(t, T) = \exp(-3\%(T - t))$

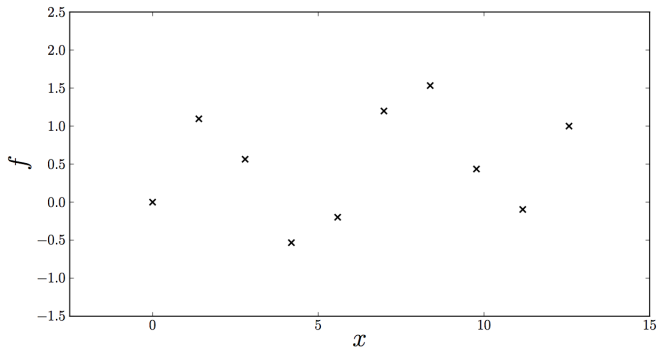


In practice, financial term-structures are constructed using deterministic interpolation techniques.

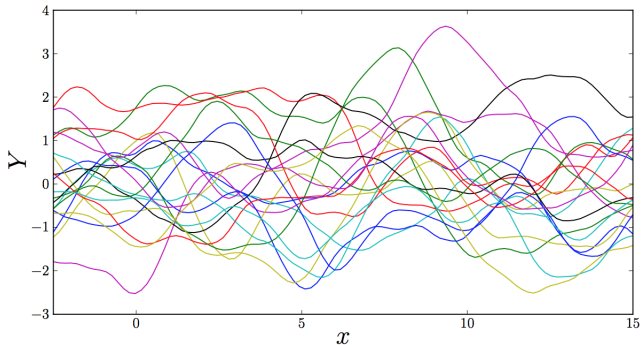
- Parametric approaches : [Nelson-Siegel](#) or [Svensson](#) models (used by most central banks)
- Non-parametric interpolation methods : shape-preserving spline techniques (lack of interpretability but better ability to fit the data).

**Could we propose an arbitrage-free interpolation method that additionally allows for quantification of uncertainty ?**

A function  $f$  is only known at a limited number of points  $x_1, \dots, x_n$



The (unknown) function  $f$  is assumed to be a sample path of a **Gaussian process**  $Y$



## Definition : Gaussian process (GP) or Gaussian random field

A Gaussian process is a collection of random variables, any finite number of which have (consistent) joint Gaussian distributions.

A Gaussian process  $(Y(x), x \in \mathbb{R}^d)$  is characterized by its **mean function**

$$\mu : x \in \mathbb{R}^d \longrightarrow \mathbb{E}(Y(x)) \in \mathbb{R}.$$

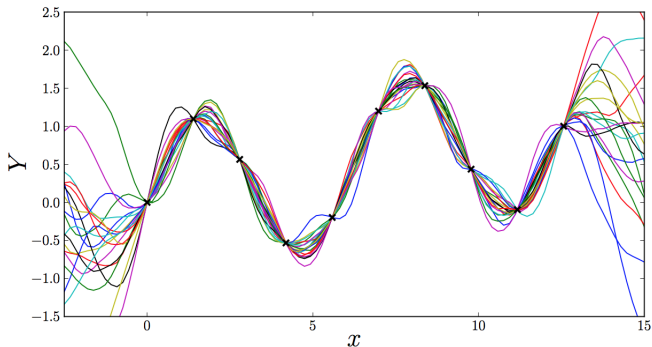
and its **covariance function**

$$K : (x, x') \in \mathbb{R}^d \times \mathbb{R}^d \longrightarrow \text{Cov}(Y(x), Y(x')) \in \mathbb{R}.$$

**Table:** Some popular covariance functions  $K(x, x')$  used in 1D kriging methods.

Name	Expression	Class
Gaussian	$\sigma^2 \exp\left(-\frac{(x-x')^2}{2\theta^2}\right)$	$\mathcal{C}^\infty$
Matérn 5/2	$\sigma^2 \left(1 + \frac{\sqrt{5} x-x' }{\theta} + \frac{5(x-x')^2}{3\theta^2}\right) \exp\left(-\frac{\sqrt{5} x-x' }{\theta}\right)$	$\mathcal{C}^2$
Matérn 3/2	$\sigma^2 \left(1 + \frac{\sqrt{3} x-x' }{\theta}\right) \exp\left(-\frac{\sqrt{3} x-x' }{\theta}\right)$	$\mathcal{C}^1$
Exponential	$\sigma^2 \exp\left(-\frac{ x-x' }{\theta}\right)$	$\mathcal{C}^0$

The estimation of  $f$  relies on the conditional distribution of  $Y$  given the observed values  $y_i = f(x_i)$  at points  $x_i$ ,  $i = 1, \dots, n$ .



- $\mathbf{X} = (x_1, \dots, x_n)^\top \in \mathbb{R}^{n \times d}$  : some design points
- $\mathbf{y} = (y_1, \dots, y_n)^\top \in \mathbb{R}^n$  : observed values of  $f$  at these points
- $Y(\mathbf{X}) = (Y(x_1), \dots, Y(x_n))^\top$  : vector composed of  $Y$  at point  $\mathbf{X}$

## The conditional process is still a Gaussian Process

Let  $Y$  be a GP with mean  $\mu$  and covariance function  $K$ . The conditional process  $Y \mid Y(\mathbf{X}) = \mathbf{y}$  is a GP with mean function

$$\eta(x) = \mu(x) + \mathbf{k}(x)^\top \mathbb{K}^{-1}(\mathbf{y} - \boldsymbol{\mu}), \quad x \in \mathbb{R}^d$$

and covariance function  $\tilde{K}$  given by

$$\tilde{K}(x, x') = K(x, x') - \mathbf{k}(x)^\top \mathbb{K}^{-1} \mathbf{k}(x'), \quad x, x' \in \mathbb{R}^d$$

where  $\boldsymbol{\mu} = \mu(\mathbf{X}) = (\mu(x_1), \dots, \mu(x_n))^\top$ ,  $\mathbb{K}$  is the covariance matrix of  $Y(\mathbf{X})$  and  $\mathbf{k}(x) = (K(x, x_1), \dots, K(x, x_n))^\top$

Recall that, in our term-structure construction problem, the (unknown) real function  $f$  satisfies some linear equality constraints of the form

$$A \cdot f(X) = \mathbf{b}, \quad (1)$$

where

- $A$  is a given matrix of dimension  $n \times m$
- $X = (x_1, \dots, x_m)^\top \in \mathbb{R}^{m \times d}$
- $f(X) = (f(x_1), \dots, f(x_m))^\top \in \mathbb{R}^m$
- $\mathbf{b} \in \mathbb{R}^n$

# Extension to linear equality constraints

- $X = (x_1, \dots, x_m)^\top \in \mathbb{R}^{m \times d}$  : some design points
- $\mathbf{b} = (b_1, \dots, b_n)^\top \in \mathbb{R}^n$  : right-hand side of the linear system
- $Y(X) = (Y(x_1), \dots, Y(x_m))$  : vector composed of  $Y$  at point  $X$

## The conditional process is still a Gaussian Process

Let  $Y$  be a GP with mean  $\mu$  and covariance function  $K$ . The conditional process  $Y \mid AY(X) = \mathbf{b}$  is a GP with mean function

$$\eta(x) = \mu(x) + (\mathbf{A}\mathbf{k}(x))^\top (\mathbf{A}\mathbb{K}\mathbf{A}^\top)^{-1} (\mathbf{b} - \mathbf{A}\mu), \quad x \in \mathbb{R}^d$$

and covariance function  $\tilde{K}$  given by

$$\tilde{K}(x, x') = K(x, x') - (\mathbf{A}\mathbf{k}(x))^\top (\mathbf{A}\mathbb{K}\mathbf{A}^\top)^{-1} \mathbf{A}\mathbf{k}(x'), \quad x, x' \in \mathbb{R}^d$$

where  $\mu = \mu(X) = (\mu(x_1), \dots, \mu(x_m))^\top$ ,  $\mathbb{K}$  is the covariance matrix of  $Y(X)$ ,  $\mathbf{k}(x) = (K(x, x_1), \dots, K(x, x_m))^\top$



**New formulation of the problem** : estimation of an unknown function  $f$  given that

$$\begin{cases} A \cdot f(X) = \mathbf{b} \\ f \in \mathcal{M} \end{cases}$$

where  $\mathcal{M}$  is the set of (say) non-increasing functions.

**Problem** : The conditional process is not a Gaussian process anymore.

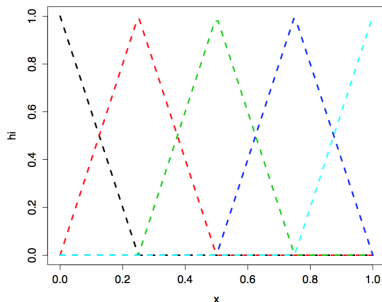
- How to cope with the **infinite-dimensional** monotonicity constraints?
- Which estimator could we propose for the term-structure?

# Extension to monotonicity constraints (1D case)

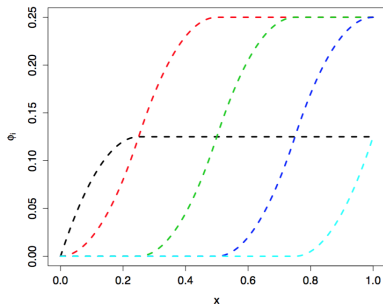
**Proposed methodology** : On an interval  $D = [\underline{x}, \bar{x}]$  of  $\mathbb{R}$ , we construct a **finite-dimensional approximation** of  $Y$  for which the monotonicity constraint is easy to check.

- Regular subdivision  $u_0 < \dots < u_N$  of  $D$  with a constant mesh  $\delta$
- Set of increasing basis functions  $(\phi_i)_{i=0, \dots, N}$  defined on this subdivision

$$h_i(x) := \max\left(1 - \frac{|x - u_i|}{\delta}, 0\right)$$



$$\phi_i(x) = \int_{\underline{x}}^x h_i(u) du$$



## Proposition (Maatouk and Bay, 2014b)

Let  $Y$  be a zero-mean GP with covariance function  $K$  and with almost surely differentiable paths.

- The finite-dimensional process  $Y^N$  defined on  $D$  by

$$Y^N(x) = Y(u_0) + \sum_{j=0}^N Y'(u_j) \phi_j(x)$$

uniformly converges to  $Y$ , almost surely.

- $Y^N$  is non-decreasing (resp. non-increasing) on  $D$  if and only if  $Y'(u_j) \geq 0$  (resp.  $Y'(u_j) \leq 0$ ) for all  $j = 0, \dots, N$ .
- Let  $\xi := (Y(u_0), Y'(u_0), \dots, Y'(u_N))^T$ , then  $\xi \sim \mathcal{N}(0, \Gamma^N)$  where

$$\Gamma^N = \begin{bmatrix} K(u_0, u_0) & \frac{\partial K}{\partial x'}(u_0, u_j) \\ \frac{\partial K}{\partial x}(u_i, u_0) & \frac{\partial^2 K}{\partial x \partial x'}(u_i, u_j) \end{bmatrix}_{0 \leq i, j \leq N}$$

# Extension to monotonicity constraints (1D case)

For a given covariance function  $K$ , we assume that the unknown function  $f$  is a sample path of the GP

$$Y^N(x) = \eta + \sum_{j=0}^N \xi_j \phi_j(x), \quad x \in D,$$

where  $\xi := (\eta, \xi_0, \dots, \xi_N)^\top \sim \mathcal{N}(0, \Gamma^N)$ .

Kriging  $f$  is equivalent to find the conditional distribution of  $Y^N$  given

$$\begin{cases} A \cdot Y^N(X) = \mathbf{b} & \text{linear equality condition} \\ \xi_j \leq 0, j = 0, \dots, N & \text{monotonicity constraint} \end{cases}$$

# Extension to monotonicity constraints (1D case)

Or equivalently, to find the distribution of the truncated Gaussian vector  $\xi \sim \mathcal{N}(0, \Gamma^N)$  given

$$\begin{cases} A \cdot \Phi \cdot \xi = \mathbf{b} & \text{linear equality condition} \\ \xi_j \leq 0, j = 0, \dots, N & \text{monotonicity constraint} \end{cases}$$

where  $\Phi$  is a  $m \times (N + 2)$  matrix defined as

$$\Phi_{i,j} := \begin{cases} 1 & \text{for } i = 1, \dots, m \text{ and } j = 1, \\ \phi_{j-2}(x_i) & \text{for } i = 1, \dots, m \text{ and } j = 2, \dots, N + 2. \end{cases}$$

Which estimator could we use for  $f$  ?

We consider the **mode of the truncated gaussian process** (most probable path) :

$$M_K^N(x | A, \mathbf{b}) = \nu + \sum_{j=0}^N \nu_j \phi_j(x),$$

where  $\nu = (\nu, \nu_0, \dots, \nu_N)^\top \in \mathbb{R}^{N+2}$  is the solution of the following convex optimization problem :

$$\nu = \arg \min_{\mathbf{c} \in \mathcal{C} \cap \mathcal{I}(A, \mathbf{b})} \left( \frac{1}{2} \mathbf{c}^\top (\Gamma^N)^{-1} \mathbf{c} \right),$$

with

- $\mathcal{C} = \{ \boldsymbol{\xi} \in \mathbb{R}^{N+2} : \xi_j \leq 0, j = 0, \dots, N \}$
- $\mathcal{I}(A, \mathbf{b}) = \{ \boldsymbol{\xi} \in \mathbb{R}^{N+2} : A \cdot \Phi \cdot \boldsymbol{\xi} = \mathbf{b} \}$

## Efficient simulation of the truncated Gaussian vector

1) Simulate a truncated vector  $\xi$  given the linear equality constraint :

$$Z \sim \{\xi \mid B \cdot \xi = \mathbf{b}\} \sim \mathcal{N}\left(\left((B\Gamma^N)^\top (B\Gamma^N B^\top)^{-1} \mathbf{b}, \Gamma^N - (B\Gamma^N)^\top (B\Gamma^N B^\top)^{-1} B\Gamma^N\right)\right)$$

where  $B = A \cdot \Phi$ .

2) Simulate

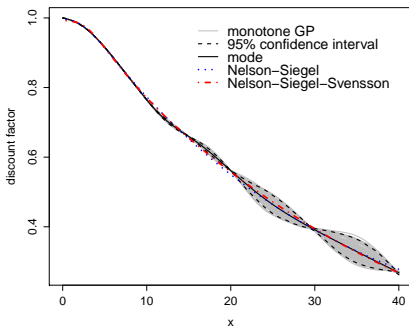
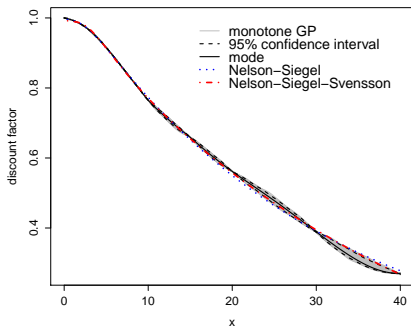
$$\{Z \mid \xi_j \leq 0, j = 0, \dots, N\} \sim \{\xi \mid B \cdot \xi = \mathbf{b} \text{ and } \xi_j \leq 0, j = 0, \dots, N\}$$

by an accelerated rejection sampling method (we use the method proposed in [Maatouk and Bay, 2014a](#))

3) The corresponding sample curves  $Y^N(\cdot) = \eta + \sum_{j=0}^N \xi_j \phi_j(\cdot)$  satisfies the constraints on the entire domain  $D$ .

# Kriging of OIS discount curves

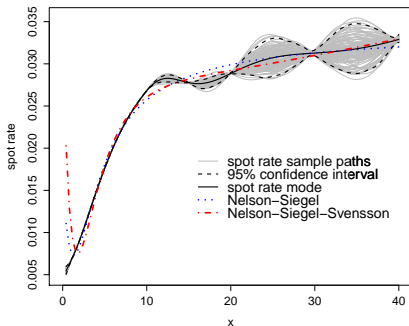
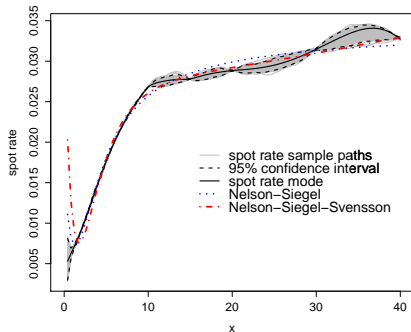
- We compare two covariance functions : **Gaussian** and **Matérn 5/2**
- Hyper-parameters  $\theta$  and  $\sigma$  are estimated using cross-validation
- Comparison with **Nelson-Siegel** and **Svensson** curve fitting



**Discount curves.**  $N = 50$ , 100 sample paths. Left : Gaussian covariance function. Right : Matérn 5/2 covariance function. OIS data of 03/06/2010.

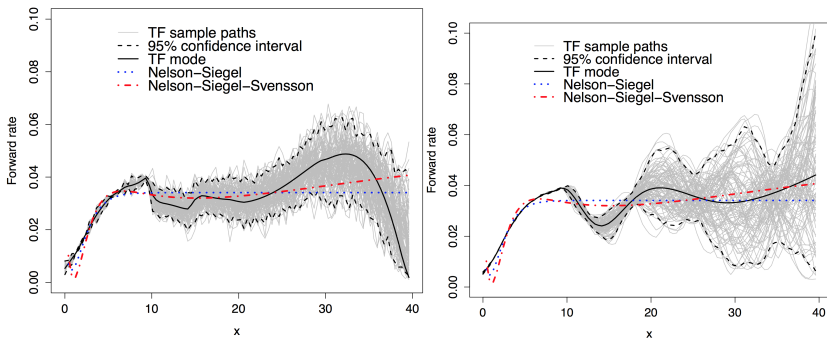


Corresponding spot rate curves :  $-\frac{1}{x} \log P(x)$



**Spot rate curves.** Left : Gaussian covariance function. Right : Matérn 5/2 covariance function. OIS data of 03/06/2010. The black solid line is the most likely spot rate curve  $-\frac{1}{x} \log M_K^N(x | A, \mathbf{b})$ .

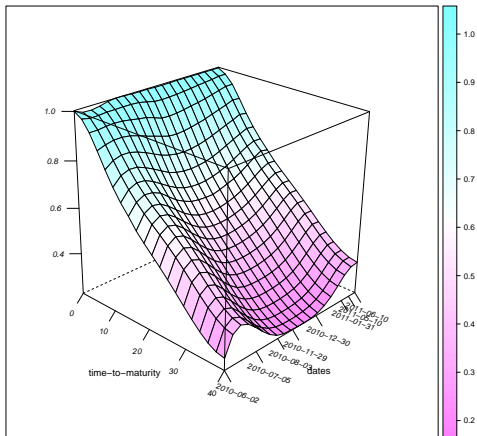
Corresponding forward rate curves :  $-\frac{d}{dx} \log P(x)$



**Spot rate curves.** Left : Gaussian covariance function. Right : Matérn 5/2 covariance function. OIS data of 03/06/2010. The black solid line is the most likely forward rate curve  $-\frac{d}{dx} \log M_K^N(x | \mathbf{A}, \mathbf{b})$ .

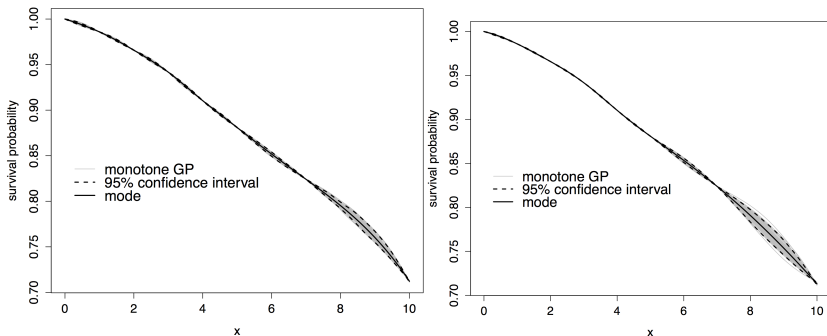
# Kriging of OIS discount curves (2D)

The previous approach can be extended in dimension 2.



**Discount curves.** OIS discount factors as a function of time-to-maturities and quotation dates.

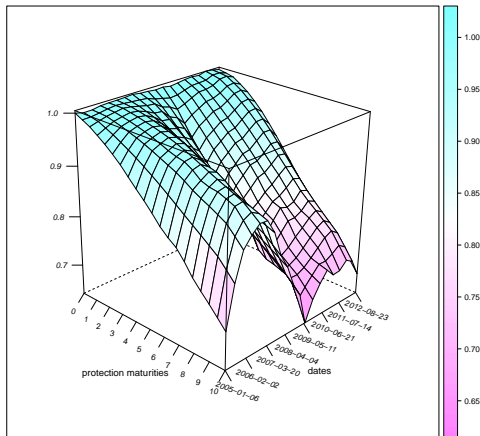
## Implied survival function of the Russian sovereign debt



**CDS implied survival curves.**  $N = 50$ , 100 sample paths. Left : Gaussian covariance function. Right : Matérn 5/2 covariance function. CDS spreads as of 06/01/2005.

# Kriging of CDS-implied default distribution (2D)

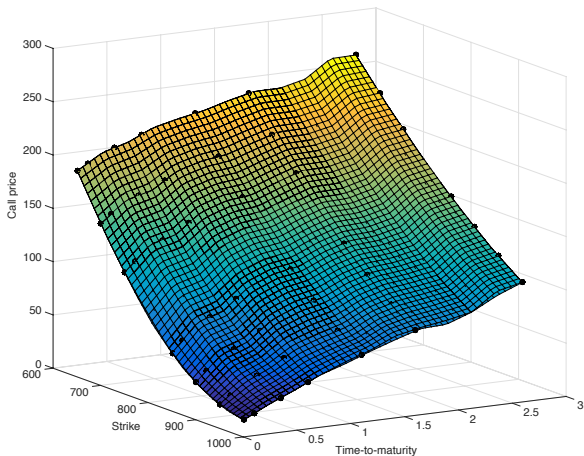
The previous approach can be extended in dimension 2.



**Survival curves.** CDS implied survival probabilities as a function of time-to-maturities and quotation dates.

- Impact of curve uncertainty on the assessment of related products and their associated hedging strategies
- What if the underlying market quotes are not reliable due to e.g. market illiquidity (data observed with a noise) ?
- Kriging of **arbitrage-free volatility surfaces** ?

# Kriging of arbitrage-free volatility surface



Thanks for your attention.





Bachoc, F. (2013).

Cross Validation and Maximum Likelihood estimations of hyper-parameters of Gaussian processes with model misspecification.

*Computational Statistics & Data Analysis*, 66(0) :55 – 69.



Bay, X., Grammont, L., and Maatouk, H. (2016).

Generalization of the Kimeldorf-Wahba correspondence for constrained interpolation.



Cousin, A., Maatouk, H., and Rullière, D. (2015).

Kriging of Financial Term-Structures.



Cousin, A. and Niang, I. (2014).

On the Range of Admissible Term-Structures.



Golchi, S., Bingham, D., Chipman, H., and Campbell, D. (2015).

Monotone Emulation of Computer Experiments.

*SIAM/ASA Journal on Uncertainty Quantification*, 3(1) :370–392.



Maatouk, H. and Bay, X. (2014a).

A New Rejection Sampling Method for Truncated Multivariate Gaussian Random Variables Restricted to Convex Sets.

*To appear in Monte Carlo and Quasi-Monte Carlo Methods 2014, Springer-Verlag, Berlin 2016.*



Maatouk, H. and Bay, X. (2014b).

Gaussian Process Emulators for Computer Experiments with Inequality Constraints.

*in revision SIAM/ASA J. Uncertainty Quantification.*